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EVIDENCE-BASED DIAGNOSTICS

Risk Factors and Screening Instruments to Predict Adverse Outcomes for Undifferentiated Older Emergency Department Patients: A Systematic Review and Meta-analysis

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Abstract

Objectives: A significant proportion of geriatric patients experience suboptimal outcomes following episodes of emergency department (ED) care. Risk stratification screening instruments exist to distinguish vulnerable subsets, but their prognostic accuracy varies. This systematic review quantifies the prognostic accuracy of individual risk factors and ED-validated screening instruments to distinguish patients more or less likely to experience short-term adverse outcomes like unanticipated ED returns, hospital readmissions, functional decline, or death.

Methods: A medical librarian and two emergency physicians conducted a medical literature search of PubMed, EMBASE, SCOPUS, CENTRAL, and ClinicalTrials.gov using numerous combinations of search terms, including emergency medical services, risk stratification, geriatric, and multiple related MeSH terms in hundreds of combinations. Two authors hand-searched relevant specialty society research abstracts. Two physicians independently reviewed all abstracts and used the revised Quality Assessment of Diagnostic Accuracy Studies instrument to assess individual study quality. When two or more qualitatively similar studies were identified, meta-analysis was conducted using Meta-DiSc software. Primary outcomes were sensitivity, specificity, positive likelihood ratio (LR+), and negative likelihood ratio (LR-) for predictors of adverse outcomes at 1 to 12 months after the ED encounters. A hypothetical test-treatment threshold analysis was constructed based on the meta-analytic summary estimate of prognostic accuracy for one outcome.

Results: A total of 7,940 unique citations were identified yielding 34 studies for inclusion in this systematic review. Studies were significantly heterogeneous in terms of country, outcomes assessed, and the timing of post-ED outcome assessments. All studies occurred in ED settings and none used published clinical decision rule derivation methodology. Individual risk factors assessed included dementia, delirium, age, dependency, malnutrition, pressure sore risk, and self-rated health. None of these risk factors significantly increased the risk of adverse outcome (LR+ range = 0.78 to 2.84). The absence of dependency reduces the risk of 1-year mortality (LR- = 0.27) and nursing home placement (LR- = 0.27). Five constructs of frailty were evaluated, but none increased or decreased the risk of adverse outcome. Three instruments were evaluated in the meta-analysis: Identification of Seniors at Risk, Triage Risk Screening Tool, and Variables Indicative of Placement Risk. None of these instruments significantly increased (LR+ range for various outcomes = 0.98 to 1.40) or decreased (LR- range = 0.53 to 1.11) the risk of adverse outcomes. The test threshold for 3-month functional decline based on the most accurate instrument was 42%, and the treatment threshold was 61%.

Conclusions: Risk stratification of geriatric adults following ED care is limited by the lack of pragmatic, accurate, and reliable instruments. Although absence of dependency reduces the risk of 1-year mortality,

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no individual risk factor, frailty construct, or risk assessment instrument accurately predicts risk of adverse outcomes in older ED patients. Existing instruments designed to risk stratify older ED patients do not accurately distinguish high- or low-risk subsets. Clinicians, educators, and policy-makers should not use these instruments as valid predictors of post-ED adverse outcomes. Future research to derive and validate feasible ED instruments to distinguish vulnerable elders should employ published decision instrument methods and examine the contributions of alternative variables, such as health literacy and dementia, which often remain clinically occult.

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Patients aged 65 years and older present to emergency departments (EDs) worldwide with increasing frequency due in part to intensified clinical and fiscal challenges inside an already frayed health care safety net.^{1–3} Clinically, providing high-quality emergency care for older adults is challenging because older patients tend to have a greater burden of comorbidities than younger patients, higher rates of serious illnesses, and frequent communication barriers. From a financial standpoint, care of the older adult is further complicated by increasing external pressure from the federal government to reduce overall costs of medical care.³ Taken together, emergency providers are being pressured to make accurate decisions to admit or discharge older patients, while concurrently seeing greater numbers of patients in less time.^{4,5} Use of screening instruments may allow point-of-care identification of geriatric patients who are at increased risk for readmission or other adverse outcomes. Although most existing instruments were derived on discharged ED patients, accurate identification of high-risk subsets while in the ED opens opportunities for emergency clinicians to alter ED or post-ED trajectories for these vulnerable patients using observation units, mobile acute care for the elderly teams, Hospital at Home models, and other home health care resources.⁵ This risk stratification allows faster and more focused use of time, testing, personnel, and resources to those most in need of these services.

Decisions to admit or discharge from the ED are generally based on the patient's risk of suffering a short-term adverse outcome. Ideally, emergency clinicians could accurately and reliably identify which older adults are at increased risk for short-term, post-ED adverse outcomes, such as preventable hospital readmissions and return ED visits, functional decline, institutionalization, and/or death. Older patients are a uniquely vulnerable population at high risk for these poor outcomes as identified by the Geriatric Emergency Medicine Residency Core Competencies and the "Geriatric Emergency Department Guidelines" (http://www.saem.org/docs/education/geri_ed_guidelines_final.pdf?sfvrsn=2).^{6–8} Recognizing the increased risk of adverse outcomes among older adults, these guidelines recommend that "All geriatric patients, regardless of the presenting complaint shall be screened (on the initial index visit, not follow-up visits) using the Identification of Seniors at Risk (ISAR) tool or a similar risk screening tool."^{9,10} If sufficiently accurate, appropriate use of these tools may allow clinicians to mend the tattered ED safety net for our vulnerable older patients. However, recent studies

from Europe and North America suggest that these risk assessment instruments lack sufficient prognostic accuracy.^{11,12} Consequently, geriatric and emergency medicine research opinion leaders list development and implementation trials of prognostic screening instruments among the highest research priorities.^{13,14}

This systematic review helps define the accuracy of existing screening tools to identify older ED patients at risk of adverse outcomes. Four recent systematic reviews evaluate the content validity and prognostic accuracy of single geriatric screening instruments such as the Triage Risk Screening Tool (TRST),^{15–18} but none of these systematic reviews evaluates all available instruments or all previously described individual risk factors for post-ED adverse outcomes. The primary objective of this systematic review is to quantify the prognostic accuracy of all individual risk factors and existing instruments for use in ED settings that are designed to identify geriatric adults at increased risk of short-term (1- to 12-month) adverse outcomes following ED visits. Adverse outcomes include return to the ED, hospital admission, functional decline, institutionalization, or death.

METHODS

Search Strategy

The design and manuscript structure of this systematic review conform to the recommendations from the Meta-analysis Of Observational Studies in Epidemiology (MOOSE) statement and Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.^{19,20} In conjunction with a medical librarian (SF), two investigators (CRC, TH) searched the medical literature from 1950 to January 2014. The medical librarian developed search strategies for the concepts of ED; people 60 years and older; screening; and adverse outcomes including functional decline, readmission, institutionalization, and death. These strategies were established using a combination of standardized terms and key words and were implemented in PubMed 1946-, Embase 1947-, Scopus 1823-, Cochrane Central Register of Controlled Trials (CENTRAL), and clinicaltrials.gov. All searches were completed in January 2014 and limited to English using database-supplied limits. To exclude animals, the librarian used the human filter for PubMed recommended in Cochrane Handbook for Systematic Reviews of Interventions²¹ as a model to create similar filters for the other databases searched. All results were exported to EndNote. We used the automatic duplicate finder in EndNote and 910 duplicates

were assumed to be accurately identified and removed, yielding a total of 7,940 unique citations. Nineteen trials were located in ClinicalTrials.gov. Full search strategies for PubMed and Embase are provided in Data Supplement S1 (available as supporting information in the online version of this paper).

Two authors (CRC, TH) reviewed the titles and abstracts to identify potentially relevant articles, which were then retrieved and the full articles reviewed. These authors then independently reviewed these articles for inclusion criteria. In addition, two authors (ES, SF) reviewed abstracts accepted for presentation at national emergency medicine conferences and published in *Academic Emergency Medicine*, *Annals of Emergency Medicine*, *Journal of the American Geriatrics Society*, and *European Geriatric Medicine* from 1990 through March 2014.

Studies and scientific research abstracts were included if they recruited a population of general geriatric adults (age ≥ 65 years) in ED settings. We sought to identify instruments that risk stratify undifferentiated older adults regardless of their presenting complaints or ED diagnoses, so disease-specific instruments such as those that risk stratify patients with congestive heart failure²² or pneumonia²³ were excluded. Similarly, instruments or risk factors for common geriatric syndromes such as standing level falls,²⁴ delirium,²⁵ and dementia²⁶ were excluded because the presence of these conditions identifies potentially vulnerable older adult patients and merits separate systematic reviews to assess each specific risk factor and instrument. Studies that reported sufficient detail on prognostic test and criterion standard results to reconstruct two-by-two tables to estimate both sensitivity and specificity were included. We contacted the authors of studies that assessed prognostic accuracy if they did not report sufficient detail to reconstruct two-by-two tables to obtain the contingency tables for our meta-analysis. If the authors responded and provided the contingency tables, then these studies were included in this systematic review. Letters or scientific abstracts with original research data were included. We excluded non-English language manuscripts, narrative reviews, case reports, and studies focused on therapy.

Individual Evidence Quality Appraisal

Two authors (CRC, TH) used the revised Quality Assessment Tool for Diagnostic Accuracy Studies (QUADAS-2) for systematic reviews to evaluate the overall quality of evidence for the identified trials.²⁷ Discrepant quality assessments were adjudicated by discussion. Statistical agreement between the two reviewers was assessed via a kappa analysis using SPSS 20. QUADAS-2 consists of nine signaling questions and four applicability questions that qualitatively assess patient selection, index test ascertainment and reproducibility, criterion standard timing and acceptability, and uniformity of obtainment and analysis for the index and criterion standard tests. The authors determined a priori to assess the quality of individual trials for the purposes of this prognostic meta-analysis by considering the following study characteristics. The ideal patient population would be those admitted or discharged from an ED with predischarge

risk assessment for short-term adverse outcomes assessed by the ED medical staff. Studies that recruited (or obtained prognostic instrument data from) a portion of patients outside of the ED were rated as low applicability. Studies that evaluated patients exclusively in the hospital or at home following ED evaluations were excluded. If a study used research personnel to administer the prognostic screening instruments rather than standard ED clinical personnel, we rated the conduct applicability as potentially low since this administration of the instrument does not reflect real-world practice. Studies that failed to explicitly blind outcome assessors to the screening instrument results or interpretation were labeled as high risk for bias.

Data Analysis

Two authors (CRC, TH) independently abstracted data from the included studies. Information abstracted included the individual study setting, inclusion criteria, mean patient age, study design, adverse outcomes assessed, outcome prevalence, and prognostic test properties. To compute meta-analytic summary estimates when more than one study assessed the same index test at the same threshold for the same or similar outcomes at the same follow-up interval, we combined eligible trials' data using Meta-DiSc (Hospital Universitario Ramón y Cajal, Madrid, Spain) using a random-effects model.²⁸ Interstudy heterogeneity was assessed with pooled estimates of sensitivity and specificity using the DerSimonian-Laird random effects model, and statistical heterogeneity was reported using the index of inconsistency (I^2).²⁹⁻³¹ Pooled estimates of dichotomous positive (LR+) and negative (LR-) likelihood ratios are also reported from the random-effects model. The level of agreement between reviewers for the QUADAS-2 assessment was quantified using a kappa analysis and the qualitative level of agreement was rated as previously described by Byrt.³² Based on the forms of bias identified by the QUADAS-2 assessment, we performed a post hoc sensitivity analysis of our results excluding all nonpublished data. Publication bias was not assessed because of the questionable validity of this approach when assessing diagnostic test meta-analyses.³³

Test-Treatment Threshold. The Pauker and Kassirer decision threshold model is based on seven variables: false-negative and false-positive proportions, sensitivity, specificity, risk of a diagnostic test, risk of treatment, and anticipated benefit of treatment.³⁴ Evidence-based estimates for each of these variables were abstracted from our meta-analysis to derive theoretical test and treatment thresholds for ED management of potentially vulnerable geriatric adults. Recognizing that these estimates are likely based on inadequate and biased research, an interactive Microsoft Excel calculator is provided with this article to permit readers to alter assumptions to recompute thresholds using different estimates of test performance or anticipated risks and benefits that may be more applicable to the end users' patient populations and clinical environments (Data Supplement S2, available as supporting information in the online version of this paper).

RESULTS

The PubMed search identified 2,682 citations, the EMBASE search identified 5,683 citations, while CENTRAL and SCOPUS yielded 170 and 80 citations, respectively. The hand search of non-peer-reviewed research abstracts yielded 55 studies. When the results were combined and the duplicates removed, 7,940 citations remained. After the titles and abstracts were reviewed to identify potentially relevant articles, which were then retrieved and the full manuscripts reviewed, 34 geriatric prognostic studies were included in this systematic review (Figure 1), including two letters, four abstracts, and 28 manuscripts. We present a detailed summary of the included trials, all of which were conducted between 1999 and 2012, in Data Supplement S3 (available as supporting information in the online version of this paper).

The authors' QUADAS-2 assessment of quality had a kappa ranging from 0.52 to 1.00 (Table 1), representing fair to excellent agreement.³² No study used a case-control design. Most of the studies excluded patients who were lost to follow-up, many of the studies had unequal numbers of patients assessed for different risk factors, and unexplained exclusions were common. The predictor variables (index test) were obtained prospectively in the weeks to months before outcomes were ascertained, so those collecting the index test were presumed to be blinded to the primary outcomes (criterion standard) unless a retrospective chart review methodology was employed. However, most studies failed to explicitly state that outcome assessors were blinded to the index tests being assessed, so there was a high risk of incorporation bias.³⁵ One study used a statistically derived definition of high hospital utilization,³⁶ and this statistical threshold might not be relevant across all health

care settings, but all other outcomes were judged to be important for emergency medicine clinicians.

Individual Risk Factors

In five studies from France, Dramé et al. and Dhaussy et al. evaluated multiple risk factors for adverse outcomes in admitted medical patients following episodes of ED care.³⁷⁻⁴¹ The outcomes assessed included mortality at 6 weeks to 2 years and institutionalization at 1 year. The variables assessed included malnutrition, age, functional limitations and dependency, presence of severe comorbidities, balance issues, number of offspring, living situation, and the presence of delirium or dementia. Malnutrition risk was evaluated using the Mini Nutritional Assessment Short Form.⁴² Functional ability was assessed using the Katz Activities of Daily Living scale,⁴³ dementia using Folstein's Mini-Mental State Examination,⁴⁴ and delirium using geriatrician application of DSM-IV criteria. Balance issues were assessed using the one leg balance test,⁴⁵ and comorbid health issues were quantified using the Charlson index.⁴⁶ Assessing prognostic accuracy using a LR+ threshold of ≥ 10 and a LR- threshold of ≤ 0.1 ,⁴⁷ none of the predictor variables increased or decreased the risk of mortality or nursing home placement enough to be clinically useful (Table 2). The highest LR+ was 2.8 (presence of severe comorbidity) and the lowest LR- was 0.27 (absence of dependency). Two other studies assessed isolated single risk factors. Chang et al.⁴⁸ assessed self-rated health stratified as bad ("fair" or "poor") or nonbad ("good" or "excellent") as a predictor of ED returns at 1 and 3 months. Minnee and Wilkinson⁴⁹ assessed living alone as a predictor of 1-week ED returns. Only one study assessed each risk factor for a given outcome, so no meta-analysis of individual predictors was possible.

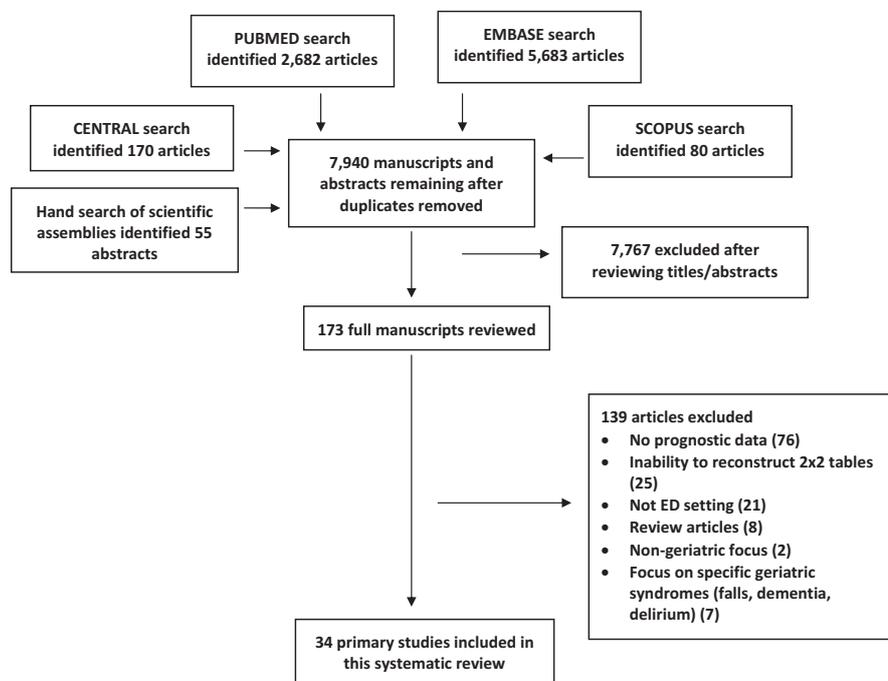


Figure 1. Study selection process.

Table 1
Overview of Quality Assessment of Included Studies

First Author, Year	Sample	Inappropriate Exclusions?	Patients and Settings Match Study Question?	Index Test Interpretation Blinded?	Index Test Conduct Applicability	Acceptable Outcome Standards?	Outcomes Assessed by Blinded Assessor?	Assessed Outcomes Pertinent?	All Patients Analyzed?
Asomaning 2014 ⁶¹	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	No	Yes	No
Bolanos 2010 ⁶²	Nonconsecutive	Yes	No	Uncertain	Uncertain	Uncertain	Uncertain	Yes	No
Braes 2009 ⁶³	Uncertain	Yes	Yes	Uncertain	No	No	Uncertain	Yes	No
Braes 2010 ⁶⁴	Consecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No
Buurman 2011 ⁶⁵	Consecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No
Carpenter 2009 ⁶⁶	Nonconsecutive	Yes	Yes	Yes	No	No	Yes	Yes	No
Carpenter 2012 ⁶⁷	Consecutive	Yes	Yes	Yes	No	No	Yes	Yes	No
Chang 2012 ⁴⁸	Consecutive	Yes	Yes	Yes	No	No	Yes	Yes	No
Conroy 2013 ⁵⁶	Nonconsecutive	Yes	No	Uncertain	No	No	Uncertain	Yes	No
Dendukuri 2004 ⁶⁸	Consecutive	No	Yes	Yes	No	No	Uncertain	Yes	Yes
Deschodt 2011 ⁶⁹	Consecutive	Yes	No	Yes	No	No	Uncertain	Yes	No
Deschodt 2012 ⁷⁰	Nonconsecutive	Uncertain	No	Yes	No	Uncertain	Uncertain	Yes	Uncertain
Dhaussy 2012 ⁴⁰	Nonconsecutive	Yes	No	Yes	No	No	Uncertain	Yes	Uncertain
Di Bari 2010 ⁵⁷	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Uncertain	Yes	Uncertain
Di Bari 2012 ⁷¹	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Yes	Yes	No
Dramé 2008 ³⁸	Nonconsecutive	Yes	No	Yes	No	No	Yes	Yes	No
Dramé 2008 ³⁷	Nonconsecutive	Yes	No	Yes	No	No	Yes	Yes	No
Dramé 2011 ³⁹	Nonconsecutive	Yes	No	Yes	No	No	Yes	Yes	No
Dramé 2011 ⁵⁴	Nonconsecutive	Yes	No	Yes	No	No	Yes	Yes	No
Dramé 2012 ⁴¹	Nonconsecutive	Yes	No	Yes	No	No	Yes	Yes	No
Edmans 2013 ⁷²	Nonconsecutive	Yes	No	Yes	No	No	Uncertain	Yes	No
Fan 2006 ⁷⁷	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Yes	Yes	No
Gentile 2013 ⁷³	Nonconsecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No

(Continued)

Table 1 (continued)

First Author, Year	Sample	Inappropriate Exclusions?	Patients and Settings Match Study Question?	Index Test Interpretation Blinded?	Index Test Conduct Applicability	Acceptable Outcome Standards?	Outcomes Assessed by Blinded Assessor?	Assessed Outcomes Pertinent?	All Patients Analyzed?
Graf 2012 ⁷⁴	Nonconsecutive	Yes	No	Yes	No	No	Uncertain	Yes	No
Graf 2012 ⁸⁰	Nonconsecutive	Yes	No	Yes	No	No	Uncertain	Yes	No
Hustey 2007 ⁷⁸	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Yes	Yes	No
Lee 2008 ⁷⁹	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Uncertain	Yes	No
McCusker 1999 ⁹	Nonconsecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No
McCusker 2000 ³⁶	Nonconsecutive	Yes	Yes	Yes	No	No	Uncertain	Uncertain	No
Meldon 2003 ¹⁰	Nonconsecutive	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Minnee 2011 ⁴⁹	Nonconsecutive	Yes	Yes	Uncertain	Yes	Uncertain	Uncertain	Yes	No
Moons 2007 ¹¹	Uncertain	No	Yes	Yes	No	No	Uncertain	Yes	Yes
Salvi 2009 ⁷⁵	Nonconsecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No
Salvi 2012 ⁷⁶	Nonconsecutive	Yes	Yes	Yes	No	No	Uncertain	Yes	No
Salvi 2012 ¹²	Nonconsecutive	Yes	Yes	Yes	Yes	No	Uncertain	Yes	No
Kappa	0.927	1.000	0.811	0.523	1.000	1.000	1.000	0.653	0.641

Dramé et al. also evaluated four validated measures of frailty (Winograd et al.⁵⁰ Rockwood et al.,⁵¹ Schoevaerds et al.,⁵² and Donini et al.⁵³) as prognostic risk factors for institutionalization or death at 1 year.⁵⁴ Stratifying the results by the highest level of frailty for each instrument, the highest LR+ for mortality was 2.9 (Rockwood⁵¹) and the lowest LR- was 0.42 (Schoevaerds et al.⁵²). In other words, the presence of “frailty” as defined by Rockwood increases the odds of death by 2.9-fold, while the absence of “frailty” as defined by Schoevaerds et al. decreases the odds of death by 0.42. However, the Schoevaerds et al. instrument identified 67% of patients as severely frail. None

of the instruments was sufficiently prognostic for 1-year nursing home placement (Table 3). In England, 645 patients in the acute medical unit over the age of 70 years were assessed with the Canadian Study on Health and Aging (CSHA) Clinical Frailty Scale⁵⁵ as a predictor of 30- or 90-day hospital readmission.⁵⁶ The CSHA frailty assessment was a similarly inaccurate predictor of hospital readmission (Table 3).

Screening Instruments

Seven different geriatric prognostic screening instruments were assessed in the included studies: Identification of Seniors at Risk (ISAR),⁹ TRST,¹⁰ the Silver

Table 2
Individual Predictors of Post-ED Adverse Outcomes

Finding, Study	Outcome Studied	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
Malnutrition risk Dramé 2008 ³⁷ MNA-SF < 12 abnormal	Mortality at 6 weeks	90 (84–95)	28 (26–31)	1.26 (1.18–1.35)	0.35 (0.20–0.58)
Presence of delirium Dramé 2008 ³⁷ Katz Activities of Daily Living Dramé 2008 ³⁷ ≥5 abnormal	Mortality at 6 weeks	33 (25–41)	81 (79–83)	1.74 (1.32–2.29)	0.83 (0.73–0.94)
Age ≥ 85 years Dramé 2011 ³⁹	Mortality at 6 weeks	80 (72–87)	52 (49–55)	1.68 (1.51–1.86)	0.38 (0.27–0.54)
Inability to toilet Dramé 2011 ³⁹	Institutionalization at 1 year	64 (59–69)	57 (54–60)	1.48 (1.34–1.65)	0.63 (0.55–0.73)
Presence of balance disorders Dramé 2011 ³⁹	Institutionalization at 1 year	76 (71–80)	47 (44–50)	1.43 (1.32–1.56)	0.51 (0.42–0.62)
Presence of dementia Dramé 2011 ³⁹	Institutionalization at 1 year	57 (52–63)	51 (48–55)	1.18 (1.06–1.32)	0.83 (0.73–0.95)
Lives alone Dramé 2012 ⁴¹	Institutionalization at 1 year	63 (57–67)	62 (58–65)	1.62 (1.45–1.82)	0.61 (0.53–0.70)
Inability to transfer Dramé 2012 ⁴¹	Institutionalization at 1 year	38 (33–43)	58 (55–61)	0.91 (0.78–1.06)	1.06 (0.97–1.17)
Severe comorbidity Dhaussy 2012 ⁴⁰ Dependency Dhaussy 2012 ⁴⁰	Mortality at 1 year	76 (71–80)	42 (39–45)	1.31 (1.21–1.42)	0.57 (0.47–0.69)
Pressure sore risk Dhaussy 2012 ⁴⁰ Malnutrition risk Dhaussy 2012 ⁴⁰	Mortality at 1 year	5 (3–7)	98 (97–99)	2.84 (1.49–5.42)	0.97 (0.95–0.99)
Delirium present Dhaussy 2012 ⁴⁰ Dementia present Dhaussy 2012 ⁴⁰	Mortality at 1 year	95 (92–97)	20 (17–23)	1.18 (1.13–1.23)	0.27 (0.18–0.41)
Lives alone Dhaussy 2012 ⁴⁰	Mortality at 1 year	59 (55–64)	70 (66–73)	1.96 (1.72–2.22)	0.58 (0.52–0.66)
<2 children Dhaussy 2012 ⁴⁰	Mortality at 1 year	87 (83–90)	33 (30–37)	1.31 (1.23–1.39)	0.39 (0.30–0.51)
Dependency Dhaussy 2012 ⁴⁰	Mortality at 1 year	26 (22–30)	83 (80–85)	1.52 (1.23–1.39)	0.89 (0.84–0.95)
Dementia present Dhaussy 2012 ⁴⁰	Mortality at 1 year	52 (48–57)	58 (55–62)	1.26 (1.11–1.41)	0.82 (0.73–0.91)
Delirium present Dhaussy 2012 ⁴⁰	Nursing home placement at 1 year	38 (33–43)	58 (55–61)	0.91 (0.78–1.06)	1.06 (0.97–1.17)
Unplanned hospital readmission Dhaussy 2012 ⁴⁰	Nursing home placement at 1 year	53 (47–58)	57 (54–61)	1.23 (1.09–1.39)	0.83 (0.73–0.93)
Severe comorbidity Dhaussy 2012 ⁴⁰	Nursing home placement at 1 year	95 (92–97)	19 (16–22)	1.17 (1.12–1.22)	0.27 (0.17–0.43)
Pressure sore risk Dhaussy 2012 ⁴⁰	Nursing home placement at 1 year	63 (57–67)	62 (58–65)	1.62 (1.45–1.82)	0.61 (0.53–0.70)
Self-rated health Chang 2012 ⁴⁸	ED returns 30 days	17 (13–21)	78 (76–81)	0.78 (0.60–1.01)	1.06 (1.00–1.12)
Self-rated health Chang 2012 ⁴⁸	ED returns 90 days	54 (48–59)	38 (34–41)	0.86 (0.76–0.97)	1.23 (1.06–1.44)
Lives alone Minnee 2011 ⁴⁹	ED returns 7 days	3 (1–5)	97 (96–98)	1.04 (0.52–2.09)	1.00 (0.98–1.02)
	Functional decline 30 days	55 (50–60)	66 (63–69)	1.61 (1.42–1.83)	0.68 (0.61–0.77)
	Functional decline 90 days	68 (49–83)	44 (36–53)	1.21 (0.92–1.59)	0.73 (0.44–1.24)
	ED returns 7 days	62 (48–75)	42 (33–51)	1.07 (0.83–1.39)	0.90 (0.60–1.35)
	ED returns 7 days	62 (51–72)	46 (35–58)	1.16 (0.90–1.50)	0.81 (0.57–1.16)
	ED returns 7 days	59 (49–69)	41 (30–52)	1.00 (0.78–1.28)	1.00 (0.70–1.43)
	ED returns 7 days	21 (5–51)	69 (51–83)	0.68 (0.22–2.08)	1.15 (0.80–1.63)

MNA-SF = Mini Nutritional Assessment Short-Form; LR = likelihood ratio.

Code,⁵⁷ Variables Indicative of Placement risk (VIP),⁵⁸ Mortality Risk Index, Rowland,⁵⁹ and Runciman.⁶⁰ The components, scoring, and interpretation for each of these instruments are provided in Data Supplement S4 (available as supporting information in the online version of this paper). The ISAR was assessed in 19 studies involving 14,440 patients.^{9,11,12,36,61–76} The TRST was assessed in 14 studies involving 7,016 patients.^{10–12,63–67,69,70,74,77–79} The VIP was assessed by four studies involving 1,765 patients.^{63,64,69,70} The Runciman instrument was assessed in three studies involving 512 patients.^{11,60,65} The Rowland instrument^{11,65} and Silver Code^{57,71} were each assessed by two studies involving 464 and 12,451 patients, respectively. The studies evaluated a wide range of outcomes at post-ED time frames ranging from 14 days to 1 year. Most studies had research teams or geriatric specialists administer the screening instruments (Data Supplement S3), but no significant prognostic accuracy difference is noted when ED personnel (nurses) administered the instruments. The outcomes assessed are detailed in Data Supplement S3 and included unanticipated ED returns, inpatient length of stay, functional decline, hospital readmission, nursing home placement, reduced mental well-being, reduced quality of life, high hospital utilization, and death. Several studies reported composite outcomes of various combinations of undesired outcomes as “any adverse outcome.”^{66,67,70,76,77,79}

The ISAR studies assessed the instrument’s prognostic accuracy for ED return visits at 30 to 180 days, functional decline at 30 to 90 days, and hospital readmission at 30 to 180 days. In addition, multiple studies reported composite outcomes of “any adverse outcome” at 30 to 180 days, although the components of the combined endpoint varied from study to study. Several studies reported prognostic accuracy estimates at various thresholds ranging from one to three “yes” responses on the ISAR.^{36,64,69,70,78} For this meta-analysis, the origi-

nal ISAR threshold for “high risk” of ≥ 2 was used across studies. The pooled estimates of sensitivity and specificity for all outcomes at all follow-up intervals demonstrated statistically significant heterogeneity with I^2 often $>50\%$ (Data Supplement S5, available as supporting information in the online version of this paper). However, pooled estimates of LR+ and LR– generally demonstrated low heterogeneity. The pooled estimates of ISAR prognostic accuracy did not differ significantly when sensitivity analysis was performed by excluding unpublished abstracts. Based upon the summary estimates of LR+ and LR–, the ISAR is not sufficiently accurate to predict increased or decreased risk of ED return visits, functional decline, hospital readmission, or any adverse outcome (Table 4).

Two groups modified the ISAR in attempts to improve the overall prognostic accuracy of the instrument.^{62,80} Graf et al.⁸⁰ made four modifications (Data Supplement S4). Item 2 was changed from “increased dependency” to “increased dependency during the past 24 hours.” They edited item 3 from “history of hospital admissions during past 6 months” to “history of hospital admissions for 1 night during 6 months.” Item 4 changed from “visual problems (with or without glasses)” to “visual problems not corrected by glasses,” while item 6 defined polypharmacy as six or more drugs rather than three. They assessed readmissions at 1 and 12 months. Bolanos et al.⁶² also changed the definition of polypharmacy to six or more drugs, and that was the only change they made to the ISAR. The modifications of the ISAR did not improve prognostic accuracy. The Graf modifications yielded a LR+ of 1.22 (95% confidence interval [CI] = 1.04 to 1.41) and 1.49 (95% CI = 1.24 to 1.79) for 1- and 12-month readmissions, as well as LR– of 0.65 (95% CI = 0.43 to 0.97) and 0.50 (95% CI = 0.37 to 0.67), respectively. Similarly, the Bolanos modifications demonstrated LR+ of 1.01 (95% CI = 0.85 to 1.20) and LR– of 0.96 (95% CI = 0.51 to

Table 3
Frailty Assessment Instruments and Post-ED Adverse Outcomes

Finding, Study	Outcome	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
Winograde Grade 3 Dramé 2011 ⁵⁴	Institutionalization at 1 year	15 (11–19)	86 (84–89)	1.10 (0.81–1.51)	0.98 (0.93–1.04)
Rockwood Grade 3 Dramé 2011 ⁵⁴	Institutionalization at 1 year	0 (0–1)	100 (99–100)	0.62 (0.07–5.50)	1.00 (0.99–1.01)
Schoevaerdt Grade 3 Dramé 2011 ⁵⁴	Institutionalization at 1 year	85 (81–89)	40 (37–43)	1.42 (1.32–1.52)	0.37 (0.29–0.48)
Donini Grade 3 Dramé 2011 ⁵⁴	Institutionalization at 1 year	76 (71–80)	32 (29–35)	1.11 (1.04–1.20)	0.76 (0.62–0.93)
Winograde Grade 3 Dramé 2011 ⁵⁴	Mortality at 1 year	21 (17–25)	89 (87–91)	1.89 (1.41–2.52)	0.89 (0.84–0.94)
Rockwood Grade 3 Dramé 2011 ⁵⁴	Mortality at 1 year	1 (0–2)	100 (99–100)	2.88 (0.48–17.16)	1.00 (0.99–1.00)
Schoevaerdt Grade 3 Dramé 2011 ⁵⁴	Mortality at 1 year	83 (79–86)	41 (37–44)	1.40 (1.30–1.50)	0.42 (0.33–0.53)
Donini Grade 3 Dramé 2011 ⁵⁴	Mortality at 1 year	75 (70–79)	32 (23–35)	1.09 (1.02–1.17)	0.80 (0.66–0.97)
CSHA Frailty Index Conroy 2013 ⁵⁶	Hospital readmission At 1 month	27 (19–35)	72 (69–76)	0.96 (0.70–1.31)	1.01 (0.91–1.14)
	At 3 months	25 (20–31)	72 (68–75)	0.89 (0.69–1.15)	1.04 (0.95–1.14)

CSHA = Canadian Study of Health and Aging; LR = likelihood ratio.

Table 4
ISAR Prognostic Accuracy

Outcome	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
30-day ED returns				
Asomaning 2013 ⁶¹	47 (24–71)	38 (31–44)	0.76 (0.47–1.23)	1.40 (0.89–2.21)
Bolanos 2010 ⁶²	86 (72–95)	10 (6–15)	0.96 (0.84–1.09)	1.40 (0.59–3.30)
Carpenter 2012 ⁶⁷	63 (47–76)	25 (15–39)	0.84 (0.64–1.10)	1.47 (0.82–2.63)
Moons 2007 ¹¹	100 (66–100)	38 (26–50)	1.52 (1.21–1.91)	0.13 (0.01–2.01)
Salvi 2012 ⁷⁶	65 (49–78)	50 (41–58)	1.28 (0.98–1.68)	0.71 (0.47–1.08)
<i>Pooled estimate</i>	<i>69 (62–76)</i>	<i>39 (36–43)</i>	<i>1.06 (0.83–1.35)</i>	<i>1.09 (0.70–1.70)</i>
90-day ED returns				
Carpenter 2009 ⁶⁶	63 (47–76)	25 (15–39)	0.84 (0.64–1.10)	1.47 (0.82–2.63)
Carpenter 2012 ⁶⁷	88 (62–98)	25 (15–39)	1.17 (0.92–1.49)	0.49 (0.12–1.94)
Moons 2007 ¹¹	80 (59–93)	40 (27–55)	1.34 (1.00–1.81)	0.50 (0.21–1.16)
<i>Pooled estimate</i>	<i>72 (61–81)</i>	<i>30 (23–38)</i>	<i>1.09 (0.83–1.43)</i>	<i>0.79 (0.34–1.84)</i>
180-day ED returns				
Di Bari 2010 ⁵⁷	50 (47–53)	69 (66–72)	1.62 (1.44–1.81)	0.72 (0.67–0.78)
Salvi 2009 ⁷⁵	73 (63–82)	43 (31–55)	1.28 (1.01–1.61)	0.63 (0.41–0.96)
Salvi 2012 ⁷⁶	61 (51–71)	59 (47–70)	1.48 (1.08–2.03)	0.66 (0.48–0.91)
Salvi 2012 ¹²	74 (71–77)	39 (36–42)	1.21 (1.14–1.29)	0.67 (0.58–0.76)
<i>Pooled estimate</i>	<i>61 (59–63)</i>	<i>51 (49–53)</i>	<i>1.38 (1.14–1.67)</i>	<i>0.71 (0.66–0.75)</i>
30-day functional decline				
Braes 2009 ⁶³	79 (63–90)	35 (27–43)	1.20 (0.99–1.47)	0.62 (0.33–1.15)
Carpenter 2012 ⁶⁷	87 (78–93)	29 (20–40)	1.23 (1.05–1.43)	0.45 (0.24–0.84)
Deschodt 2011 ⁶⁹	88 (84–92)	19 (16–23)	1.09 (1.03–1.16)	0.61 (0.42–0.89)
Salvi 2009 ⁷⁵	78 (62–89)	47 (37–57)	1.45 (1.13–1.86)	0.48 (0.26–0.89)
<i>Pooled estimate</i>	<i>86 (83–89)</i>	<i>27 (24–30)</i>	<i>1.19 (1.07–1.34)</i>	<i>0.56 (0.43–0.72)</i>
90-day functional decline				
Braes 2009 ⁶³	74 (57–87)	36 (28–45)	1.15 (0.92–1.45)	0.73 (0.41–1.30)
Carpenter 2009 ⁶⁶	77 (64–88)	36 (23–51)	1.21 (0.94–1.56)	0.63 (0.34–1.17)
Carpenter 2012 ⁶⁷	86 (71–95)	33 (18–52)	1.30 (0.99–1.70)	0.41 (0.16–1.05)
Edmans 2013 ⁷²	79 (70–87)	38 (33–43)	1.28 (1.13–1.46)	0.54 (0.36–0.82)
<i>Pooled estimate</i>	<i>79 (73–84)</i>	<i>37 (33–41)</i>	<i>1.25 (1.14–1.38)</i>	<i>0.53 (0.44–0.77)</i>
30-day hospital readmission				
Braes 2010 ⁶⁴	61 (41–78)	30 (23–38)	0.86 (0.63–1.18)	1.32 (0.79–2.23)
Deschodt 2012 ⁷⁰	86 (76–93)	18 (15–22)	1.05 (0.95–1.16)	0.77 (0.42–1.41)
Graf 2012 ⁷⁴	92 (84–97)	22 (17–27)	1.18 (1.07–1.29)	0.38 (0.18–0.79)
<i>Pooled estimate</i>	<i>85 (79–90)</i>	<i>21 (18–24)</i>	<i>1.08 (0.94–1.23)</i>	<i>0.75 (0.37–1.56)</i>
90-day hospital readmission				
Braes 2010 ⁶⁴	70 (54–83)	33 (25–42)	1.04 (0.82–1.31)	0.92 (0.55–1.55)
Edmans 2013 ⁷²	76 (69–82)	33 (29–37)	1.14 (1.02–1.26)	0.72 (0.54–0.97)
Graf 2012 ⁷⁴	93 (87–97)	28 (22–35)	1.30 (1.18–1.43)	0.24 (0.12–0.47)
<i>Pooled estimate</i>	<i>82 (77–86)</i>	<i>32 (29–35)</i>	<i>1.18 (1.05–1.34)</i>	<i>0.57 (0.30–1.10)</i>
180-day hospital readmission				
Di Bari 2010 ⁵⁷	80 (76–83)	29 (26–32)	1.12 (1.06–1.19)	0.70 (0.58–0.85)
Graf 2012 ⁷⁴	91 (86–94)	32 (25–41)	1.34 (1.19–1.52)	0.28 (0.17–0.47)
Salvi 2012 ¹²	77 (74–80)	38 (35–41)	1.24 (1.17–1.32)	0.60 (0.51–0.71)
<i>Pooled estimate</i>	<i>80 (78–82)</i>	<i>34 (32–36)</i>	<i>1.22 (1.11–1.34)</i>	<i>0.54 (0.39–0.75)</i>
30-day any adverse outcome				
Carpenter 2012 ⁶⁷	92 (75–99)	35 (21–51)	1.42 (1.11–1.81)	0.22 (0.05–0.89)
Deschodt 2012 ⁷⁰	90 (83–95)	18 (15–22)	1.10 (1.03–1.18)	0.54 (0.30–0.97)
Salvi 2012 ⁷⁶	70 (57–80)	50 (41–59)	1.38 (1.09–1.74)	0.61 (0.41–0.91)
<i>Pooled estimate</i>	<i>84 (78–88)</i>	<i>25 (22–28)</i>	<i>1.26 (1.03–1.55)</i>	<i>0.56 (0.40–0.77)</i>
90-day any adverse outcome				
Carpenter 2009 ⁶⁶	85 (75–92)	30 (15–49)	1.21 (0.94–1.56)	0.50 (0.23–1.09)
Carpenter 2012 ⁶⁷	88 (73–96)	37 (19–58)	1.39 (1.02–1.90)	0.34 (0.13–0.88)
Edmans 2013 ⁷²	71 (66–75)	43 (34–52)	1.24 (1.05–1.45)	0.68 (0.53–0.88)
<i>Pooled estimate</i>	<i>74 (70–78)</i>	<i>40 (33–47)</i>	<i>1.25 (1.11–1.42)</i>	<i>0.60 (0.44–0.83)</i>
180-day any adverse outcome				
Di Bari 2010 ⁵⁷	72 (68–76)	25 (22–28)	0.96 (0.91–1.02)	1.11 (0.94–1.31)
McCusker 1999 ⁹	72 (68–76)	58 (55–61)	1.71 (1.57–1.87)	0.48 (0.42–0.56)
Salvi 2012 ⁷⁶	68 (59–76)	61 (48–72)	1.72 (1.25–2.38)	0.53 (0.39–0.72)
<i>Pooled estimate</i>	<i>72 (69–74)</i>	<i>44 (42–46)</i>	<i>1.40 (0.88–2.24)</i>	<i>0.66 (0.37–1.19)</i>

ISAR = Identification of Seniors at Risk.

1.81). Graf also derived a multiple regression model based on four risk factors (Data Supplement S4) with LR+ of 1.1 and 1.1 and LR– of 0.15 (95% CI = 0 to 0.97) and 0.25 (95% CI = 0.1 to 0.67) for readmission at 1 and 12 months, respectively.⁸⁰

The TRST studies assessed the instrument's prognostic accuracy for ED returns at 30 to 120 days, functional decline at 30 to 90 days, and hospital readmission at 30 to 180 days. One study reported the prognostic accuracy of TRST for functional decline with the outcome

defined by either activities of daily living or instrumental activities of daily living, but the prognostic accuracy did not differ substantially, regardless of which outcome definition was used.⁷⁸ Additionally, five studies assessed a composite outcome of “any adverse outcome” at 30 days using various combinations of endpoints.^{10,67,70,77,79} Two studies evaluated “any adverse outcome” at 90 days^{65,66} and another two at 120 days.^{77,79} Several studies reported prognostic accuracy estimates at various thresholds ranging from one to three “yes” responses on the TRST.^{10,64,69,70,79} For this meta-analysis, the original threshold of ≥ 2 was used across studies. The pooled estimates of sensitivity and specificity for all outcomes at all follow-up intervals demonstrated statistically significant heterogeneity with I^2 often $>50\%$ (Data Supplement S6, available as supporting information in the online version of this paper). Pooled estimates of LR+ and LR- again generally demonstrated low heterogeneity. The pooled estimates of TRST prognostic accuracy did not differ significantly when sensitivity analysis was performed by excluding unpublished abstracts. Based upon the summary estimates of LR+ and LR- ratios, the TRST is not sufficiently accurate to predict increased or decreased risk of ED returns, functional decline, hospital readmission, or any adverse outcome (Table 5).

Two studies assessed the VIP instrument’s prognostic accuracy for hospital readmission or functional decline at 30 days.^{64,70} For this meta-analysis, we used the threshold of ≥ 1 to define “abnormal” across studies. The pooled estimates of sensitivity and specificity for all outcomes at all follow-up intervals demonstrated variable statistical heterogeneity with I^2 ranging from 0% to 99.5% (Data Supplement S7, available as supporting information in the online version of this paper). Based on the summary estimates of LR+ and LR-, the VIP is not sufficiently accurate to predict increased or decreased risk of either post-ED hospital readmission or functional decline (Table 6).

Five studies reported sufficient detail to reconstruct 2×2 tables at various thresholds (≥ 1 , ≥ 2 , and ≥ 3) for ISAR, TRST, and VIP (Table 7).^{10,37,64,69,73} Assessing prognostic accuracy for these instruments at different thresholds did not improve on the summary estimates of accuracy from our meta-analysis. In addition, one trial reported sufficient detail to compute interval likelihood ratios (iLR).^{64,81} For the ISAR and the outcome of 14-day readmissions, the $iLR_{1-2} = 1.68$ and $iLR_{2-3} = 0.94$, whereas for the TRST $iLR_{1-2} = 1.03$ and $iLR_{2-3} = 0.61$, and for the VIP $iLR_{1-2} = 1.09$ and $iLR_{2-3} = 1.00$. The prognostic accuracy of these three instruments does not improve using either different thresholds to define abnormal or by using iLR.

Several instruments were evaluated in single trials so meta-analysis was not performed for these (Table 8). In addition to the ISAR and TRST, Buurman et al.⁶⁵ evaluated the Rowland and Runciman instruments for 6-month ED revisits, readmissions, mortality, or any combination of these three outcomes, in Belgium. Neither instrument demonstrated significant prognostic accuracy to increase or decrease the risk of vulnerability for these outcomes. In Italy, Di Bari et al.^{57,71} derived and evaluated the “Silver Code” in two trials. In the derivation trial, they assessed 1-year mortality using two

thresholds of abnormal: ≥ 4 or ≥ 11 .⁵⁷ The derivation and validation sets were consistent, but neither threshold was sufficiently prognostic for 1-year mortality. Di Bari et al.⁷¹ subsequently evaluated the Silver Code for 6-month outcomes of ED revisits, readmission, mortality, or any combination of these outcomes at the same two thresholds. Again, the Silver Code lacked sufficient prognostic accuracy to increase (LR+ ranges = 1.59 to 2.47) or decrease (LR- ranges = 0.70 to 0.88) the risk of these outcomes. Dramé et al.³⁸ derived the “Mortality Risk Index” (Data Supplement S4) in France to evaluate for 2-year post-ED mortality among hospitalized geriatric patients. Using either of two thresholds of abnormal (≥ 3 or ≥ 5), the Mortality Risk Index lacked prognostic accuracy to increase (LR+ = 1.41 to 1.93) or decrease (LR- = 0.34 to 0.59) the 2-year mortality risk.

DISCUSSION

Adverse outcomes after an episode of ED care occur in one-third of discharged older patients.⁸²⁻⁸⁴ This problem is not unique to the United States, with similar reports in Italy,⁸⁵ Hong Kong,⁸⁶ and Australia.⁸⁷ While several case management models advocate for evidence-based screening and engagement of nursing leadership to enhance older adult outcomes,⁸⁸ these principles depend on a valid and reliable risk stratification protocol.⁸⁹ Unfortunately, existing research on risk factors for unsatisfactory outcomes following an older adult ED evaluation fails to provide compelling evidence to justify recommending any screening instrument or specific geriatric case finding method.⁹⁰ Nonetheless, medical educators and guideline developers advocate for focused, evidence-based screening efforts to optimize geriatric outcomes.^{6,7} Because the number of geriatric EDs is expanding⁹¹ with concurrent efforts to quantify benefit for these geriatric EDs,⁹² the demand to derive and validate more accurate risk stratification instruments will continue.

None of the individual predictors of vulnerability or published risk stratification instruments demonstrate sufficient prognostic accuracy to distinguish high-risk or low-risk subsets of geriatric patients in EDs. This systematic review included five validated measures of frailty, which some might label as risk stratification instruments or distinct geriatric syndromes that merit a separate systematic review. Our a priori exclusion criteria included the geriatric syndromes of falls, delirium, and dementia, but not frailty for two reasons. First, in the seminal essay about geriatric syndromes, Inouye et al.⁹³ distinguished frailty from falls and delirium. Second, unlike falls, delirium, and dementia, no well-accepted definition or criterion standard for frailty exists. However, the concept of frailty is quite valid. By including the prognostic accuracy of frailty (which was not reported in the original studies but was provided by the investigators), future studies of ED vulnerability will be able to incorporate these constructs of frailty among the constellation of intrinsic and extrinsic factors that place some patients at increased risk.

Several recent systematic reviews have assessed one or more of these instruments in EDs. The systematic review of the TRST by Cousins et al.¹⁸ included six

Table 5
TRST Prognostic Accuracy

Outcome	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
30-day ED returns				
Carpenter 2012 ⁶⁷	80 (44–97)	47 (35–60)	1.51 (1.03–2.22)	0.43 (0.12–1.51)
Fan 2006 ⁷⁷	63 (35–85)	52 (42–62)	1.31 (0.85–2.02)	0.72 (0.37–1.38)
Moons 2007 ¹¹	67 (35–90)	47 (34–60)	1.25 (0.79–1.99)	0.71 (0.31–1.65)
<i>Pooled estimate</i>	<i>68 (51–82)</i>	<i>49 (43–56)</i>	<i>1.06 (0.83–1.35)</i>	<i>1.09 (0.70–1.70)</i>
90-day ED returns				
Carpenter 2009 ⁶⁶	63 (47–76)	38 (25–52)	1.01 (0.75–1.37)	0.98 (0.60–1.61)
Carpenter 2012 ⁶⁷	64 (31–89)	47 (33–62)	1.20 (0.72–2.01)	0.77 (0.34–1.78)
Moons 2007 ¹¹	64 (43–82)	48 (34–62)	1.23 (0.83–1.83)	0.75 (0.41–1.36)
<i>Pooled estimate</i>	<i>63 (52–73)</i>	<i>44 (36–52)</i>	<i>1.11 (0.89–1.38)</i>	<i>0.86 (0.61–1.22)</i>
120-day ED returns				
Buurman 2011 ⁶⁵	79 (66–88)	33 (28–38)	1.17 (1.00–1.37)	0.65 (0.39–1.10)
Fan 2006 ⁷⁷	60 (42–76)	54 (43–65)	1.31 (0.92–1.88)	0.74 (0.47–1.16)
<i>Pooled estimate</i>	<i>71 (61–80)</i>	<i>37 (32–42)</i>	<i>1.19 (1.03–1.38)</i>	<i>0.70 (0.50–0.98)</i>
30-day functional decline				
Braes 2009 ⁶³	79 (63–90)	47 (38–55)	1.48 (1.18–1.84)	0.46 (0.25–0.84)
Carpenter 2012 ⁶⁷	70 (51–84)	53 (38–69)	1.50 (1.01–2.22)	0.57 (0.31–1.02)
Deschodt 2011 ⁶⁹	78 (73–83)	30 (25–34)	1.11 (1.02–1.21)	0.74 (0.56–0.96)
Hustey 2007 ⁷⁸	63 (50–75)	60 (56–64)	1.58 (1.28–1.97)	0.61 (0.44–0.86)
<i>Pooled estimate</i>	<i>75 (71–79)</i>	<i>27 (24–30)</i>	<i>1.37 (1.10–1.71)</i>	<i>0.65 (0.54–0.78)</i>
90-day functional decline				
Braes 2009 ⁶³	79 (63–90)	50 (41–59)	1.58 (1.24–2.00)	0.42 (0.22–0.80)
Carpenter 2009 ⁶⁶	60 (46–74)	36 (23–51)	0.94 (0.70–1.28)	1.10 (0.67–1.81)
Carpenter 2012 ⁶⁷	61 (42–78)	50 (31–69)	1.23 (0.78–1.93)	0.77 (0.44–1.37)
<i>Pooled estimate</i>	<i>66 (57–75)</i>	<i>47 (40–54)</i>	<i>1.23 (0.87–1.75)</i>	<i>0.73 (0.42–1.27)</i>
30-day hospital readmission				
Braes 2010 ⁶³	57 (37–76)	39 (32–48)	0.94 (0.67–1.33)	1.09 (0.68–1.74)
Deschodt 2012 ⁷⁰	71 (59–81)	26 (22–30)	0.95 (0.82–1.12)	1.13 (0.77–1.67)
Fan 2006 ⁷⁷	75 (35–97)	52 (43–62)	1.57 (1.01–2.45)	0.48 (0.14–1.61)
Graf 2012 ⁷⁴	87 (78–93)	22 (17–27)	1.11 (1.00–1.23)	0.60 (0.33–1.09)
<i>Pooled estimate</i>	<i>76 (70–82)</i>	<i>30 (27–32)</i>	<i>1.06 (0.92–1.24)</i>	<i>0.90 (0.63–1.29)</i>
90-day hospital readmission				
Braes 2010 ⁶⁴	67 (51–81)	45 (36–54)	1.22 (0.94–1.59)	0.73 (0.45–1.17)
Graf 2012 ⁷⁴	88 (81–93)	24 (18–30)	1.16 (1.05–1.27)	0.51 (0.30–0.85)
<i>Pooled estimate</i>	<i>83 (76–88)</i>	<i>32 (27–37)</i>	<i>1.16 (1.06–1.28)</i>	<i>0.62 (0.43–0.85)</i>
180-day hospital readmission				
Graf 2012 ⁷⁴	89 (83–93)	27 (21–35)	1.22 (1.10–1.36)	0.41 (0.25–0.66)
Salvi 2012 ¹²	72 (68–76)	41 (38–44)	1.22 (1.14–1.30)	0.68 (0.59–0.79)
<i>Pooled estimate</i>	<i>76 (73–79)</i>	<i>40 (37–42)</i>	<i>1.22 (1.16–1.29)</i>	<i>0.56 (0.34–0.91)</i>
30-day any adverse outcome				
Carpenter 2012 ⁶⁷	69 (52–84)	55 (38–71)	1.54 (1.03–2.31)	0.56 (0.32–0.98)
Deschodt 2012 ⁷⁰	77 (68–85)	26 (22–30)	1.04 (0.93–1.16)	0.88 (0.61–1.28)
Fan 2006 ⁷⁷	65 (38–86)	52 (42–62)	1.36 (0.91–2.04)	0.67 (0.34–1.31)
Lee 2008 ⁷⁹	82 (74–88)	24 (21–28)	1.07 (0.98–1.17)	0.76 (0.53–1.10)
Meldon 2003 ¹⁰	64 (57–72)	63 (59–67)	1.74 (1.48–2.05)	0.57 (0.46–0.70)
<i>Pooled estimate</i>	<i>73 (69–77)</i>	<i>37 (35–40)</i>	<i>1.29 (1.03–1.62)</i>	<i>0.67 (0.55–0.81)</i>
90-day any adverse outcome				
Carpenter 2009 ⁶⁶	64 (52–75)	33 (17–53)	0.97 (0.71–1.31)	1.07 (0.59–1.93)
Carpenter 2012 ⁶⁷	61 (42–77)	48 (28–69)	1.17 (0.73–1.86)	0.82 (0.46–1.48)
<i>Pooled estimate</i>	<i>63 (53–72)</i>	<i>40 (27–54)</i>	<i>1.02 (0.79–1.32)</i>	<i>0.94 (0.62–1.42)</i>
120-day any adverse outcome				
Fan 2006 ⁷⁷	62 (45–77)	55 (43–66)	1.37 (0.97–1.93)	0.70 (0.45–1.09)
Lee 2008 ⁷⁹	56 (49–62)	58 (54–62)	1.33 (1.15–1.55)	0.76 (0.65–0.89)
<i>Pooled estimate</i>	<i>57 (51–63)</i>	<i>58 (54–62)</i>	<i>1.34 (1.17–1.53)</i>	<i>0.75 (0.65–0.87)</i>

TRST = Triage Risk Screening Tool.

studies and adverse outcomes as the outcome. Their conclusions mirrored ours: “the TRST is limited in its ability to discriminate between older adults with or without adverse outcome following ED discharge.” The systematic review by Beaton and Grimmer¹⁶ evaluated only functional decline and no meta-analysis is reported. They did identify two instruments that we did not include in our systematic review: the Brief Risk Identification for Geriatric Health Tool (BRIGHT)⁹⁴ and the Simplified PROFUNCTION index.⁹⁵ We excluded the

BRIGHT because the validation trial assessed point-of-care instrumental activities of daily living dependence rather than post-ED functional decline. We excluded the Simplified PROFUNCTION index because it has not been evaluated in ED settings. Bissett et al.¹⁷ assessed the psychometric properties of the ISAR and TRST, as well as two longer instruments, the Older Adult Resources and Services⁹⁶ and the Functional Status Assessment of Seniors in Emergency Departments.⁹⁷ They concluded that the shorter instruments were more

Table 6
Variables Indicative of Placement Risk Prognostic Accuracy

Outcome, Study	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
30-day hospital readmission				
Braes 2010 ⁶⁴	73 (52–88)	35 (27–43)	1.12 (0.86–1.45)	0.77 (0.40–1.51)
Deschodt 2012 ⁷⁰	81 (70–89)	13 (10–16)	0.93 (0.82–1.04)	1.48 (0.88–2.49)
<i>Pooled estimate</i>	79 (69–86)	18 (15–21)	0.98 (0.83–1.17)	1.11 (0.59–2.09)
30-day functional decline				
Braes 2010 ⁶⁴	43 (28–59)	88 (81–93)	3.55 (2.02–6.26)	0.65 (0.50–0.85)
Deschodt 2012 ⁷⁰	88 (83–910)	21 (17–25)	1.11 (1.04–1.18)	0.58 (0.41–0.84)
<i>Pooled estimate</i>	82 (77–86)	37 (33–42)	1.92 (0.58–6.41)	0.63 (0.50–0.78)

Table 7
Prognostic Accuracy at Different Thresholds

Instrument	Study	Outcome	Sensitivity	Specificity	Positive LR	Negative LR
ISAR						
≥1	Braes 2010 ⁶⁴	Hospital readmit at 30 days	79	10	0.88	2.05
≥3	Braes 2010 ⁶⁴	Hospital readmit at 30 days	43	56	0.98	1.29
≥3	Gentile 2013 ⁷³	Hospital readmit at 30 days	83	34	1.26	0.51
≥3	McCusker 2000 ³⁶	High hospital utilizer	50	74	1.96	0.67
≥3	McCusker 2000 ³⁶	High hospital utilizer	23	88	1.94	0.87
TRST						
≥1	Braes 2010 ⁶³	Hospital readmit at 30 days	75	15	0.88	1.70
≥3	Braes 2010 ⁶³	Hospital readmit at 30 days	21	81	1.13	0.97
≥1	Meldon 2003 ¹⁰	Adverse outcome at 30 days	85	32	1.25	0.47
≥3	Deschodt 2011 ⁶⁹	Functional decline at 30 days	51	59	1.23	0.84
VIP						
≥1	Braes 2010 ⁶³	Hospital readmit at 30 days	68	35	1.04	0.92
≥3	Braes 2010 ⁶³	Hospital readmit at 30 days	4	97	1.16	0.99

ISAR = Identification of Seniors at Risk; LR = likelihood ratio; TRST = Triage Risk Screening Tool; VIP = Variables Indicative of Placement risk

appropriate for timely screening, whereas the longer instruments provide a more comprehensive understanding of functional performance. They did not quantify prognostic accuracy.

The lack of an accurate screening instrument represents one barrier to more routine geriatric prognosis screening.⁹⁸ The ideal prognostic screening instrument would be well calibrated across a broad range of illness severity, disability, socioeconomic, and health literacy strata.⁹⁹ To be clinically useful, the instrument would be sufficiently accurate to significantly reduce (LR− ≤ 0.1) or increase (LR+ ≥ 10) the risk of an adverse outcome so that finite resources can be focused on appropriately high-risk individuals.⁴⁷ Using these parameters of efficacy, neither the ISAR, the TRST, nor the VIP is sufficiently accurate to be clinically useful, as demonstrated in Figure 2.

The complexity in assessing the risk of unexpected and suboptimal outcomes, such as short-term functional decline, return to the ED, hospitalization, nursing home placement, or death, exists at the level of the patient, the community, and the national health care system. First, chronological age is not synonymous with biological age, yet all of these studies evaluate the geriatric adult population as one homogenous group. In the future, genetic or tissue biomarkers might assist clinicians in distinguishing the frail elderly phenotype from

the acutely ill but otherwise healthy geriatric adult with adequate physiologic reserve.^{100–103} In addition, assessing disease-specific prognostic risk (the heart failure prognostic instruments,²² for example) in conjunction with tools like the ISAR or TRST might yield more accurate estimates of risk for short-term adverse outcomes. Second, in many patients a return to the ED or subsequent hospitalization can be expected and appropriate, but most studies did not distinguish anticipated from unexpected readmissions. A planned readmission is a repeat hospitalization following an index hospital stay that can be foreseen at the time of discharge after the index admission (delayed surgery, for example). An unplanned readmission cannot be foreseen at the time of discharge and may occur on a nonemergent or emergent basis.¹⁰⁴ Third, various geriatric syndromes complicate the post-ED management of the vulnerable elder, including delirium,¹⁰⁵ standing-level falls,¹⁰⁶ and dementia.¹⁰⁷ For example, if assessed in non-medicated, non-critically ill older adults, researchers note abnormal cognitive function in 35% of older ED patients. However, this is only recognized by emergency physicians in 6% of cases.¹⁰⁷

The next layer of complexity is at the community level. Substantial proportions of young and old patients lack adequate literacy and numeracy skills to reliably comprehend medical information in the acute health

Table 8
Other Instruments: Runciman, Rowland, and Silver Code Prognostic Accuracy

Outcome, Study	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
Mortality Risk Index Dramé 2008 ³⁸				
Mortality at 2-years				
≥ 3 abnormal	87 (84–90)	38 (35–42)	1.41 (1.32–1.50)	0.34 (0.27–0.43)
≥ 5 abnormal	59 (54–63)	70 (66–73)	1.93 (1.69–2.20)	0.59 (0.53–0.66)
Rowland				
Buurman 2011 ⁶⁵				
6-month				
ED returns	23 (13–36)	82 (77–86)	1.28 (0.75–2.17)	0.94 (0.81–1.09)
Readmission	23 (14–35)	83 (78–87)	1.35 (0.81–2.24)	0.93 (0.80–1.07)
Mortality	27 (6–61)	82 (78–86)	1.51 (0.56–4.05)	0.89 (0.62–1.28)
Any of above	25 (16–37)	83 (78–87)	1.49 (0.94–2.36)	0.90 (0.78–1.04)
Runciman				
Buurman 2011 ⁶⁵				
6-month:				
ED returns	86 (74–94)	12 (9–16)	0.97 (0.87–1.09)	1.19 (0.59–2.41)
Readmission	85 (74–92)	12 (9–16)	0.96 (0.86–1.08)	1.28 (0.67–2.43)
Mortality	82 (48–98)	12 (9–16)	0.93 (0.70–1.23)	1.53 (0.42–5.52)
Any of above	87 (77–93)	12 (9–16)	0.99 (0.89–1.09)	1.10 (0.57–2.12)
Silver Code				
Di Bari 2010 ⁵⁷				
1-year mortality				
Derivation set				
Threshold ≥ 4	84 (83–85)	33 (32–34)	1.25 (1.22–1.28)	0.49 (0.45–0.53)
Threshold ≥ 11	33 (32–35)	83 (83–84)	2.0 (1.87–2.14)	0.80 (0.78–0.82)
Validation set				
Threshold ≥ 4	84 (82–86)	32 (30–34)	1.23 (1.20–1.27)	0.50 (0.45–0.57)
Threshold ≥ 11	32 (30–34)	83 (82–84)	1.93 (1.75–2.13)	0.81 (0.79–0.84)
Silver Code				
Di Bari 2012 ⁷¹				
Threshold ≥ 4				
6-month:				
ED returns	74 (71–78)	35 (32–38)	1.15 (1.07–1.23)	0.73 (0.62–0.85)
Readmission	77 (73–81)	35 (32–38)	1.19 (1.12–1.27)	0.65 (0.54–0.78)
Mortality	90 (85–93)	35 (32–38)	1.37 (1.30–1.46)	0.30 (0.21–0.44)
Any of above	75 (72–78)	39 (35–42)	1.22 (1.14–1.31)	0.65 (0.56–0.75)
Threshold ≥ 11				
6-month:				
ED returns	27 (24–31)	84 (82–87)	1.74 (1.42–2.11)	0.86 (0.82–0.91)
Readmission	28 (24–32)	83 (80–85)	1.59 (1.31–1.93)	0.88 (0.83–0.93)
Mortality	42 (36–48)	83 (81–85)	2.47 (2.04–2.99)	0.70 (0.63–0.78)
Any of above	27 (24–30)	87 (84–89)	2.06 (1.65–2.57)	0.84 (0.80–0.88)

An 82-year-old community-dwelling and functionally independent widow presents to the ED for evaluation of unilateral hip pain. After a careful, geriatric-appropriate history and physical exam reveals no evidence of cognitive dysfunction, abuse/neglect, standing-level falls, or inappropriate polypharmacy, you obtain x-rays of the painful hip and pelvis. No radiographic pathology is identified and the pain is completely relieved with acetaminophen so you contemplate discharge home, but wonder what the risk of an unanticipated ED return is within 3 months. Based on the studies included in this meta-analysis, the 3-month ED return risk ranged from 22% to 35%, so you use a pretest probability of 35%. The most accurate prognostic instrument was the TRST (LR+ = 1.11, LR- = 0.86 from Table 5). Converting the probability to odds using the formula [odds = probability ÷ (1 - probability)] yields an odds of 0.54 [0.35 ÷ (1 - 0.35) = 0.54]. Multiplying these odds by the LR's yields 0.54 × 1.11 = 0.592 posttest odds for an individual labeled as "high risk" by TRST and 0.54 × 0.86 = 0.46 for an individual labeled as "low risk" by TRST. These posttest odds are converted back to probability using the formula [probability = odds ÷ (odds + 1)]. For example, the posttest probability of 3-month ED returns for a "high-risk" TRST patient is 0.592/1.592 = 0.372 or 37.2%. Remember that before the TRST score was used the baseline probability was 35% so using the TRST only increased the risk by 2.2%.

Figure 2. Applying these results in the clinical setting.

care setting.^{108,109} The result is miscomprehension of diagnoses and misunderstanding of discharge instructions, follow-up care plans, and prescription management.¹⁰⁸ Smaller EDs with limited access to social workers or a geriatric unit also have higher rates of return ED visits by older adults.¹¹⁰ In addition, different communities provide variable levels of access to primary care or specialists for post-ED management.¹¹¹ Therefore, comparing the prognostic accuracy of one instrument across health care settings may be pointless. A patient in a setting where every citizen has health insurance and a single, identifiable primary care physician is a patient aware of whom to see in follow-up, confident that such services will be reimbursed. On the other hand, in a pay-for-service system, a fixed-income older adult without a single primary care physician lacks such ease of access. The latter patient might need to return to the ED for the same or new medical issues or have no follow-up for months, yet the return to the ED would be attributed to the patient's vulnerability rather than the health care system's limitations. Therefore, future prognostic instruments might need to differ according to a region's health care system.

Nurse and physician educators need to recognize the limitations of existing instruments as predictors of adverse outcomes following an episode of ED care. Additionally, we dissuade researchers, payers, and policy-makers from using any of these instruments as a basis for meaningful risk stratification. Designing expensive, time-consuming interventions or protocols around these instruments is illogical. The challenge then is how to risk stratify these complex patients and where to focus future efforts at education, research, and clinical policy? One response is to reemphasize the need for ED-based geriatric research to more fully illuminate evidence-based prognosis, diagnosis, and therapy decisions.^{13,14} Another response is to explore prognostic instruments previously validated in non-ED settings to assess thus far untested risk factors.

Test–Treatment Threshold. Several ED-based studies have assessed geriatric prognostic screening instruments in quantitative, randomized trials and assessed outcomes such as hospital readmission, ED returns, outpatient referral patterns and service use, and functional decline. Caplan et al.¹¹² assessed Australian ED-based comprehensive geriatric assessment (CGA) with the intervention group receiving risk-stratified, home-based multidisciplinary assistance for up to 28 days following ED discharge. They demonstrated a 5.7% absolute risk reduction in 30-day hospital admission and a 9.9% absolute risk reduction in 18-month ED returns, but demonstrated no effect on mortality or nursing home admission. Mion et al.¹¹³ also assessed ED-based CGA in the United States and noted a 2.3% absolute risk reduction in nursing home admissions at 30 days, but no effect on overall service use at 30 or 120 days. These studies represent the “gold standard” to which more pragmatic, ED-based geriatric screening interventions aspire, but require availability of a high-quality home health care service and skilled personnel with the time to administer the lengthy CGA in the ED. Therefore, we did not use

these estimates of treatment benefit in our test–treatment computations.

Investigators in Canada used a randomized design to evaluate ISAR screening followed by nurse assessment of high-risk geriatric patients, notification of the primary care physician, and recommended home health care referrals. They demonstrated a 9.85% absolute risk reduction in 4-month functional decline, but no effect on patient depression, caregiver outcomes, or satisfaction with care.¹¹⁴ In addition, ISAR-screened patients were more likely to be referred to community wellness resources, their primary care physicians, and home care services. Unfortunately, they were also 8.8% more likely to return to the ED.¹¹⁵ Nonetheless, the ISAR screening model reduced the total cost of post-ED care over the subsequent 4 months.¹¹⁶ Therefore, we use the estimate of 9.85% reduction in 4-month functional decline in our test–treatment computations.

In addition to estimates of treatment benefit in those with “disease” and prognostic instrument accuracy, the Pauker-Kassirer test–treatment assumption variables include estimated risks of the “test” and risk of treatment in those without “disease.”³⁴ We use the prognostic accuracy of ISAR to predict 3-month functional decline since our outcome benefit estimate from the Canadian randomized trial is prevention of 4-month functional decline. However, the risk associated with administering the ISAR in EDs has not been formally assessed or quantified. Theoretical risks might include delayed diagnosis of time-dependent emergencies due to the time spent collecting the ISAR variables, increased patient angst related to an “abnormal” screening test result or other patient harm related to the downstream effect of further testing and referrals initiated due to ISAR screening results. However, we believe that these risks are quite small and use an estimate of 0.0001 for the test–treatment equation. One risk of treatment in those without “disease” is increased referral to the ED from home health services or primary care physicians. Although such referrals are probably indicated in the majority of cases, these ED returns would not otherwise have occurred without the ISAR-based intervention. These ED returns represent a measurable risk to patients, including exposure to ED treatment and infectious agents, transportation to and from the ED, and possibly hospital admission. The Canadian data suggest that the risk of increased ED returns at 1 month is 8.8%. For our test–treatment calculations we assume that this risk remains constant (8.8%) at 3 months. We acknowledge that this is an unlikely assumption, but no randomized trial data exists for 3-month ED returns following ISAR-based interventions.

These summary estimates were used to compute the test and treatment thresholds (Figure 3). The test threshold is 42% and the treatment threshold is 61%. In other words, when contemplating whether to develop an ED-based geriatric screening program, the investment in personnel training and time to screen is more likely to cause harm or lack direct patient benefit if the pretest risk of functional decline is less than 42% or greater than 61%. In the former case, the evidence suggests searching for other non-ISAR-related risks for functional decline. In the latter case, when the risk

These methods should be employed in the development, derivation, validation, and impact analysis of the next generation of prognostic instruments.¹²⁴ For example, before deriving the instrument, patients and providers should identify outcomes and potential predictors. Clinical decision rule methods also provide a tiered approach to instrument derivation, validation, and impact assessment. Barriers to implementation of standardized ED geriatric screening instruments include lack of resources and inadequate follow-up, misunderstanding the distinction between screening and assessment tools, and the need to adapt instruments to the local context.¹²⁶ In designing for implementation from the first stages of derivation, future investigators should contemplate the role of disruptive innovation in user-friendly, ED clinician-feasible instruments.¹²⁷ For example, one could use pad-based or smart phone technology to develop instruments that risk stratify patients in parallel with ongoing ED operations so as not to increase length of stay or overburden nurses and clinicians. Electronic medical record data are already being used to identify subsets of patients at increased risk for hospital readmission, and similar methods could be used to assess predischarge patient risk in ED settings in real time.^{128,129}

Finally, ED researchers should view the failures detailed above as a call to use more standardized collaborative methodology and uniform outcome markers. This would enable between-study comparisons and application of findings in real-world settings. Decades of effort by dedicated geriatric ED researchers have thus far failed to yield meaningful risk stratification tools to guide treatment for vulnerable older adults. A geriatric ED research consortium to combine talents and resources would provide an essential infrastructure to develop the instruments we need for effective risk stratification of ED geriatric patients.⁹¹

LIMITATIONS

This meta-analysis has several limitations. First, the meta-analysis of individual studies demonstrates significant statistical heterogeneity, even when assessing the same instrument for the same outcomes on similar patient populations. This heterogeneity is partially due to inconsistent definitions for outcomes and variable methods of measuring and/or obtaining the outcome measures. In addition, some studies recruited patients solely from the ED, while others included ED patients after admission. Several articles lacked sufficient details to reconstruct 2×2 tables or to ensure that similar research strategies were used, so authors were contacted when necessary and the authors' responses are included in the reported study details, QUADAS-2 assessment, and prognostic accuracy data estimates. Nonetheless, the QUADAS-2 assessment indicates several forms of potential bias, including spectrum bias and incorporation bias, since outcome assessors sometimes lacked blinding to the index test results.³⁵ The failure to use STARD methods¹²⁵ and clinical decision instrument research design strategies¹²⁴ probably also resulted in lower research quality and greater inter-study heterogeneity.

Another limitation is that a lack of sufficiently similar prognostic studies existed to perform meta-analysis for some of the instruments and outcomes, but a systematic review can only analyze previously published research. We did include unpublished research abstracts, but our a priori sensitivity analysis excluding the unpublished studies from the meta-analysis did not significantly alter our estimates of prognostic accuracy. In addition, our evaluation included several different outcomes (ED returns, hospital readmissions, functional decline, nursing home admission, death). The relative importance of each outcome for patients, clinicians, and society is undefined, but likely not equal. Lacking any empiric basis by which to weight different outcomes, we report the prognostic accuracy estimate for each as if all are equally important.

Finally, multiple unmeasured and usually clinically unrecognized confounding variables at the patient and community levels exist across studies, such as cognitive impairment, limited health literacy, fixed finances, and access to primary care including transportation. These confounding variables likely limit our ability to apply results across heterogeneous health care settings and bias estimates of prognostic accuracy for these instruments. Most of the instruments and individual risk factors assess intrinsic features of the patient, but future studies should assess these hidden variables as the search for a more accurate risk stratification instrument continues.

CONCLUSIONS

Accurate and reliable identification of vulnerable geriatric adults during episodes of ED care is essential to focus finite resources on those most likely to benefit. Existing instruments designed to risk stratify older ED patients do not accurately distinguish high- or low-risk subsets and should not be used by key stakeholders for this purpose. Additional research is required using accepted decision instrument methods while remaining cognizant of long-term implementation in heterogeneous health care settings. Current educators and researchers should emphasize the limitations of existing instruments and individual risk factors if they disseminate these predictors. Funding organizations must support additional research if the strength of evidence-based geriatric risk screening is to improve.

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Supporting Information

The following supporting information is available in the online version of this paper:

Data Supplement S1. Detailed electronic search strategy.

Data Supplement S2. Excel Test-Treatment Threshold Calculator.

Data Supplement S3. Summary of included studies.

Data Supplement S4. Components of the instruments to risk-stratify older adults in the emergency department for short-term adverse outcomes.

Data Supplement S5. Forest plots for the Identification of Seniors at Risk.

Data Supplement S6. Forest plots for the Triage Risk Screening Tool.

Data Supplement S7. Forest plots for the Variables Indicative of Placement Risk.